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Factorial experiments

n two-level factors

$$x_i = 1, -1, i = 1, 2, \dots, n$$

$$f(x_1, x_2, \dots, x_n) = \mu + \sum_{i=1}^n \beta_i x_i + \sum_{1 \leq i < j \leq n} \beta_{ij} x_i x_j + \sum_{1 \leq i < j < k \leq n} \beta_{ijk} x_i x_j x_k + \dots + \beta_{12\dots n} x_1 \dots x_n$$

μ : mean, β_i : main effect, β_{ij} : two-factor interaction, ...

full factorial: run size = 2^n

Fractional factorial designs

2^{n-p} **regular** fractional factorial designs

$n = \# \text{ of factors}$

$N = 2^{n-p} = \text{run size}$

$\frac{1}{2^p}$ -fraction of 2^n complete factorial

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Fractional Factorial Designs (FFD)

	A	B	C	a	b	c	ab	ac	bc	abc
1	-1	-1	-1	-1	-1	-1	1	1	1	-1
2	-1	-1	1	-1	-1	1	1	-1	-1	1
3	-1	1	-1	-1	1	-1	-1	1	-1	1
4	-1	1	1	-1	1	1	-1	-1	1	-1
5	1	-1	-1	1	-1	-1	-1	-1	1	1
6	1	-1	1	1	-1	1	-1	1	-1	-1
7	1	1	-1	1	1	-1	1	-1	-1	-1
8	1	1	1	1	1	1	1	1	1	1

	A	B	C	a	b	c	ab	ac	bc	abc
2	-1	-1	1	-1	-1	1	1	-1	-1	1
3	-1	1	-1	-1	1	-1	-1	1	-1	1
5	1	-1	-1	1	-1	-1	-1	-1	1	1
8	1	1	1	1	1	1	1	1	1	1

- $\text{effect}^2 = I$

$$a^2 = b^2 = c^2 = (ab)^2 = (ac)^2$$

$$= (bc)^2 = (abc)^2 = I$$

- defining relation

$$I = abc$$

- effect aliasing

$$a = bc$$

$$b = ac$$

$$c = ab$$

- un-aliased effects are orthogonal

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- A design of size N can accommodate at most $N - 1$ two-level factors
- A **saturated design** of size $N = 2^k$ can be constructed by writing all possible combinations of k factors in a $2^k \times k$ array, and then completing all possible component-wise products of the columns.
- A **regular** design with n factors is obtained by choosing n columns from the saturated design.

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If we use capital letters to denote the factors, then we also use combinations of these letters to denote interactions

A: main effect of factor A

AB: interaction of factors A and B

BCE: interaction of factors B, C and E

etc.

$$x_4 = x_1 x_2, x_5 = x_1 x_3$$

$$1 = x_1 x_2 x_4 = x_1 x_3 x_5 = x_2 x_3 x_4 x_5 \quad \text{Defining relation}$$

$$I = ABD = ACE = BCDE$$

Defining words

Defining contrasts, Defining effects

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$$1 = x_1 x_2 x_4 = x_1 x_3 x_5 = x_2 x_3 x_4 x_5$$

$$x_1 = x_2 x_4 = x_3 x_5 = x_1 x_2 x_3 x_4 x_5$$

$$A = BD = CE = ABCDE$$

In the model matrix for the full model, the columns corresponding to the main effect A and interactions BD, CE, ABCDE are identical. Therefore they are completely mixed up. We say they are **aliases** of one another.

$\{A, BD, CE, ABCDE\}$: alias set

One can estimate only one effect in each alias set, assuming that all the other effects in the same alias set are negligible.

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7 alias sets

if E is a plot factor, we would
 - remove interaction between plot & treatment
 - $E = AC = BCD$ is called confounding.

$$\boxed{A} = BD = \cancel{CE} = \cancel{ABCDE}$$

$$\boxed{B} = AD = \cancel{ACE} = \cancel{CDE}$$

$$\boxed{C} = ABCD = \cancel{AE} = \cancel{BDE}$$

$$D = \boxed{AB} = \cancel{ACDE} = \cancel{BCE}$$

$$\boxed{E} = \cancel{ABDE} = \boxed{AC} = BCD$$

$$\boxed{BC} = ACD = \cancel{ABE} = \cancel{DE}$$

$$CD = \boxed{ABC} = \cancel{ADE} = \cancel{BE}$$

In general, among the $2^n - 1$ factorial effects, $2^p - 1$ appear in the defining relation. The rest are divided into $2^{n-p} - 1$ alias sets, each of size 2^p .

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Regular fractional factorial designs have simple alias structures: any two factorial effects are either orthogonal or completely aliased.

Nonregular designs have complex alias structures that are difficult to disentangle.

Under the design defined by $I = ABD = ACE = BCDE$, the usual estimate of the main effect of A actually estimates

$A + BD + CE + ABCDE$. This is an unbiased estimate of A if all its aliases are negligible.

→ joint effect.

When some contrasts are found significant but cannot be attributed to specific effects, one has to perform follow-up experiments to resolve the ambiguity. This is called **de-aliasing**.

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Which of the following two 2^{7-2} resolution IV designs is better?

$d_1: I = DEFG = ABCDF = ABCEG$

$d_2: I = ABCF = ADEG = BCDEFG$

Minimum aberration (Fries and Hunter, 1980): *Technometrics*

Sequentially minimize A_1, A_2, \dots , where

A_i = number of words of length i in the defining relation

Word length pattern (A_1, A_2, \dots)

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Design Key

- Patterson (1965) *J. Agric. Sci.*
- Patterson (1976) *JRSS, Ser. B*
- Bailey, Gilchrist and Patterson (1977) *Biometrika*
- Bailey (1977) *Biometrika*
- Patterson and Bailey (1978) *Applied Statistics*

Factorial design construction, identification of effect aliasing and confounding with block factors

① notation & definition

T_1, \dots, T_m : m treatment factors, each has t_i levels

P_1, \dots, P_n : n plot factors (block factor, row/column factor
nested factor)

each has P_j levels.

- pseudo-factor: for a factor with $R_1^{U_1} R_2^{U_2} \dots R_r^{U_r}$ levels,
where R_1, \dots, R_r are ^{different} prime numbers, we can use

U_1 R_1 -level, U_2 R_2 -level, ..., U_r R_r -level pseudo-factors
to represent it. e.g.

4-level	X	X_1	X_2	$X: 3\text{-main effects}$
" 2-level	$0 \leftarrow 0$	0		
	$1 \leftarrow 0$	1		
	$2 \leftarrow 1$	0		
	$3 \leftarrow 1$	1		

$\boxed{X_1, X_2, X_1 X_2}$

- When a particular treatment effect is confounded with a plot effect, the plot effect is called plot alias.
- A design key specifies the plot aliases of the treatment main effects

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Situation 1: Symmetric \Rightarrow # of levels for factors are p^{m_i} , p : a prime number.

a. single replicate (i.e., $N=t$)

(i) full factorial design for treatment

(ii) fractional factorial design for treatment

b. multiple replicates (i.e. $N=2t, 3t, \dots$)

(i) full factorial for treatment

(ii) fractional factorial for treatment

Situation 2: asymmetric \Rightarrow # of levels = $R_1^{m_1} \cdot \dots \cdot R_r^{m_r}$

Q: how to find $W_T^{(i)}$ in which stratum?

Step 1. identify strata

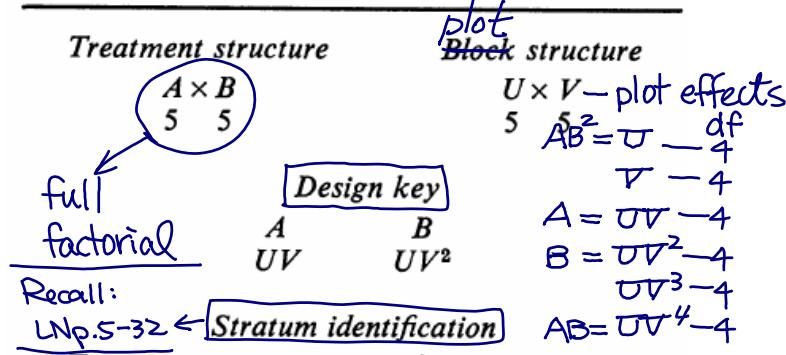
Step 2. use pseudo-factors to represent true factors

Step 3. decompose each stratum into plot effects represented by the plot pseudo-factors.

Step 4. choose a design key for the plot aliases of treatment main effects represented by pseudo-factors

Step 5. use design key to find the plot aliases of other treatment effects.

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$N=t$ A Graeco-Latin squareInverse key

U	V
$A^2 B^4$	$A^4 B$

Rules of construction

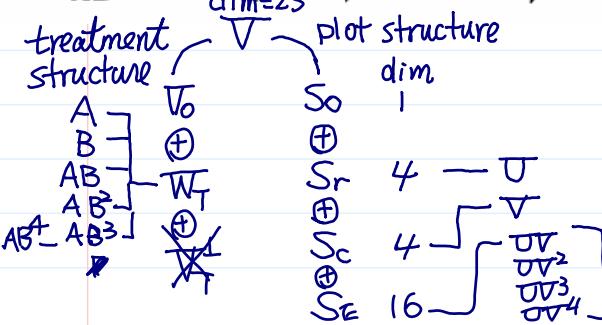
$$q(A) = q(U) + q(V) \pmod{5}$$

$$q(B) = q(U) + 2q(V) \pmod{5}$$

Construction of designColumn (level of V)

	0	1	2	3	4
Row (level of U)	00	12	24	31	43
1	11	23	30	42	04
2	22	34	41	03	10
3	33	40	02	14	21
4	44	01	13	20	32

Treatment effect	d.f.	Plot alias	Stratum
A	4	UV	UV
B	4	UV^2	UV
AB	4	$U^2 V^3$	UV
AB^2	4	U^3	U
AB^3	4	$U^4 V^2$	UV
AB^4	4	V^4	V

 $\dim=25$ 

$$AB = (UV)(UV^2) = U^2 V^3 \approx UV^4$$

$$AB^2 = UV \cdot U^2 V^4 \quad \leftarrow U^6 V^9 \pmod{5}$$

$$= U^3 V^5 = U^3 \approx U$$

$$AB^3 = \dots$$

$$AB^4 = \dots$$

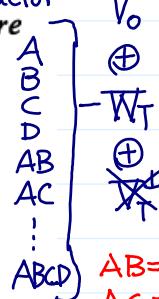
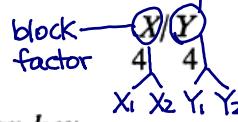
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 $N=t$ A single replicate block designTreatment structure

$A \times B \times C \times D$
2 2 2 2

full factorialDesign key

A	B	C	D
Y_1	Y_2	$X_1 \quad Y_1 \quad Y_2$	$X_2 \quad Y_1 \quad Y_2$

plot Block structure nested factor

$$AB = Y_1 Y_2$$

$$AC = X_1 Y_2$$

$$ABC = Y_1 Y_2 X_1 Y_1 Y_2 = X_1$$

$$ABD = X_2$$

$$CD = X_1 X_2$$

*assume some higher-order interactions are negligible, some df. are left for testing

Stratum $X \quad XY$ Inverse key $X_1 \quad ABC$ $X_2 \quad ABD$ $Y_1 \quad A$ $Y_2 \quad B$ Stratum contentsd.f.

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Treatment effects

ABC, ABD, CD

The rest

Partial inverse key $X_1 \quad ABC$ $X_2 \quad ABD$ Rules of construction

$$q(A) = q(Y_1) \pmod{2}$$

$$q(B) = q(Y_2) \pmod{2}$$

$$q(C) = q(X_1) + q(Y_1) + q(Y_2) \pmod{2}$$

$$q(D) = q(X_2) + q(Y_1) + q(Y_2) \pmod{2}$$

main effects AB
confounded
with plot
main effects

$A \quad B \quad C \quad D$	X_1	X_2	$\rightarrow X$
0 0 0 0	0	0	0
0 0 0 1	0	1	1
0 0 1 0	1	0	2
0 0 1 1	1	1	3
0 1 0 0	0	0	0
0 1 0 1	0	1	2
0 1 1 0	1	0	3
0 1 1 1	1	1	1
1 0 0 0	0	0	0
1 0 0 1	0	1	2
1 0 1 0	1	0	3
1 0 1 1	1	1	1
1 1 0 0	0	0	0
1 1 0 1	0	1	2
1 1 1 0	1	0	3
1 1 1 1	1	1	1

a list of treatment effects whose plot aliases are plot main effects

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