

6-1

Factorial experiments

n two-level factors

$$x_i = 1, -1, i = 1, 2, \dots, n$$

$$\begin{aligned}
 f(x_1, x_2, \dots, x_n) = & \mu + \sum_{i=1}^n \beta_i x_i + \sum_{1 \leq i < j \leq n} \beta_{ij} x_i x_j \\
 & + \sum_{1 \leq i < j < k \leq n} \beta_{ijk} x_i x_j x_k + \dots + \beta_{12 \dots n} x_1 \dots x_n
 \end{aligned}$$

μ : mean, β_i : main effect, β_{ij} : two-factor interaction,...

full factorial: run size = 2^n

Fractional factorial designs

2^{n-p} **regular** fractional factorial designs

n = # of factors

$N = 2^{n-p}$ = run size

$\frac{1}{2^p}$ -fraction of 2^n complete factorial

6-2

Fractional Factorial Designs (FFD)

	A	B	C	a	b	c	ab	ac	bc	abc
1	-1	-1	-1	-1	-1	-1	1	1	1	-1
2	-1	-1	1	-1	-1	1	1	-1	-1	1
3	-1	1	-1	-1	1	-1	-1	1	-1	1
4	-1	1	1	-1	1	1	-1	-1	1	-1
5	1	-1	-1	1	-1	-1	-1	-1	1	1
6	1	-1	1	1	-1	1	-1	1	-1	-1
7	1	1	-1	1	1	-1	1	-1	-1	-1
8	1	1	1	1	1	1	1	1	1	1

• $\text{effect}^2 = I$

$$a^2 = b^2 = c^2 = (ab)^2 = (ac)^2$$

$$= (bc)^2 = (abc)^2 = I$$

• defining relation

$$I = abc$$

• effect aliasing

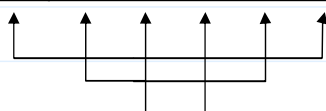
$$a = bc$$

$$b = ac$$

$$c = ab$$

• un-aliased effects are orthogonal

	A	B	C	a	b	c	ab	ac	bc	abc
2	-1	-1	1	-1	-1	1	1	-1	-1	1
3	-1	1	-1	-1	1	-1	-1	1	-1	1
5	1	-1	-1	1	-1	-1	-1	-1	1	1
8	1	1	1	1	1	1	1	1	1	1



6-3

- A design of size N can accommodate at most $N - 1$ two-level factors
- A **saturated design** of size $N = 2^k$ can be constructed by writing all possible combinations of k factors in a $2^k \times k$ array, and then completing all possible component-wise products of the columns.
- A **regular** design with n factors is obtained by choosing n columns from the saturated design.

6-4

If we use capital letters to denote the factors, then we also use combinations of these letters to denote interactions

A: main effect of factor A

AB: interaction of factors A and B

BCE: interaction of factors B, C and E

etc.

$$x_4 = x_1x_2, x_5 = x_1x_3$$

$$1 = x_1x_2x_4 = x_1x_3x_5 = x_2x_3x_4x_5 \quad \text{Defining relation}$$

$$I = ABD = ACE = BCDE$$

Defining words

Defining contrasts, Defining effects

6-5

$$1 = x_1x_2x_4 = x_1x_3x_5 = x_2x_3x_4x_5$$

$$x_1 = x_2x_4 = x_3x_5 = x_1x_2x_3x_4x_5$$

$$A = BD = CE = ABCDE$$

In the model matrix for the full model, the columns corresponding to the main effect A and interactions BD, CE, ABCDE are identical. Therefore they are completely mixed up. We say they are **aliases** of one another.

$$\{A, BD, CE, ABCDE\}: \text{alias set}$$

One can estimate only one effect in each alias set, assuming that all the other effects in the same alias set are negligible.

6-6

7 alias sets

of E is a plot factor, we would
 - remove interaction between plot & treatment
 - $E = AC = BCD$ is called confounding.

$$\begin{aligned}
 A &= BD = \cancel{CE} = \cancel{ABCDE} \\
 B &= AD = \cancel{ABCE} = \cancel{CDE} \\
 C &= ABCD = \cancel{AE} = \cancel{BDE} \\
 D &= \cancel{AB} = \cancel{ACDE} = \cancel{BCE} \\
 E &= \cancel{ABDE} = \cancel{AC} = BCD \\
 BC &= ACD = \cancel{ABE} = \cancel{DE} \\
 CD &= \cancel{ABC} = \cancel{ADE} = \cancel{BE}
 \end{aligned}$$

In general, among the $2^n - 1$ factorial effects, $2^p - 1$ appear in the defining relation. The rest are divided into $2^{n-p} - 1$ alias sets, each of size 2^p .

6-7

Regular fractional factorial designs have simple alias structures: any two factorial effects are either orthogonal or completely aliased.

Nonregular designs have complex alias structures that are difficult to disentangle.

Under the design defined by $I = ABD = ACE = BCDE$, the usual estimate of the main effect of A actually estimates $A + BD + CE + ABCDE$. This is an unbiased estimate of A if all its aliases are negligible.

→ joint effect.

When some contrasts are found significant but cannot be attributed to specific effects, one has to perform follow-up experiments to resolve the ambiguity. This is called **de-aliasing**.

6-8

Which of the following two 2^{7-2} resolution IV designs is better?

$$d_1: I = DEFG = ABCDF = ABCEG$$

$$d_2: I = ABCF = ADEG = BCDEFG$$

Minimum aberration (Fries and Hunter, 1980): Technometrics

Sequentially minimize A_1, A_2, \dots , where
 A_i = number of words of length i in the defining relation

Word length pattern (A_1, A_2, \dots)

6-9

Design Key

- Patterson (1965) *J. Agric. Sci.*
 - Patterson (1976) *JRSS, Ser. B*
 - Bailey, Gilchrist and Patterson (1977) *Biometrika*
 - Bailey (1977) *Biometrika*
 - ◆ Patterson and Bailey (1978) *Applied Statistics*
- Factorial design construction, identification of effect aliasing and confounding with block factors

6-10

notation & definition

T_1, \dots, T_m : m treatment factors, each has t_i levels

P_1, \dots, P_n : n plot factors (block factor, row/column factor, nested factor)
each has p_j levels.

- pseudo-factor: for a factor with $p_1^{v_1} p_2^{v_2} \dots p_r^{v_r}$ levels, where p_1, \dots, p_r are ^{different} prime numbers, we can use v_1 p_1 -level, v_2 p_2 -level, \dots , v_r p_r -level pseudo-factors to represent it. e.g.

4-level	X	X_1	X_2	X: 3-main effects ↓ $X_1, X_2, X_1 X_2$
2-level	0	0	0	
	1	0	1	
	2	1	0	
	3	1	1	

- When a particular treatment effect is confounded with a plot effect, the plot effect is called plot alias.
- A design key specifies the plot aliases of the treatment main effects

6-11

Situation 1: Symmetric \Rightarrow # of levels for factors are p^{m_i} , p : a prime number.

a. single replicate (i.e., $N=t$)

- full factorial design for treatment
- fractional factorial design for treatment

b. multiple replicates (i.e., $N=2t, 3t, \dots$)

- full factorial for treatment
- fractional factorial for treatment

Situation 2: asymmetric \Rightarrow # of levels = $p_1^{m_1} \dots p_r^{m_r}$

Q: how to find $W_T^{(i)} \in$ which stratum?

step 1. identify strata

step 2. use pseudo-factors to represent true factors

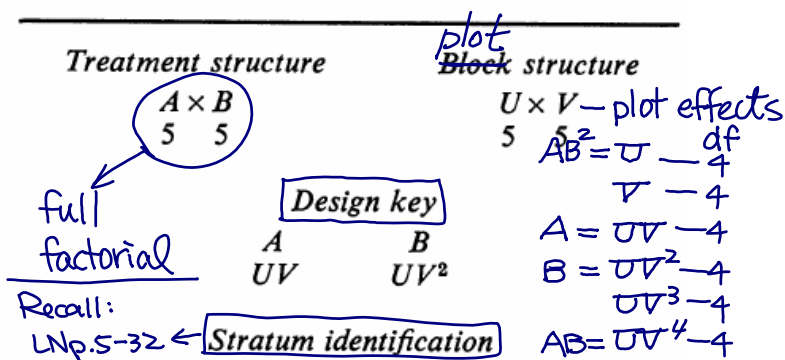
step 3. decompose each stratum into plot effects represented by the plot pseudo-factors.

step 4. choose a design key for the plot aliases of treatment main effects represented by pseudo-factors

step 5. use design key to find the plot aliases of other treatment effects.

6-12

$N=t$ A Graeco-Latin square



Inverse key

$$\begin{array}{cc} U & V \\ A^2 B^4 & A^4 B \end{array}$$

Rules of construction

$$\begin{aligned} q(A) &= q(U) + q(V) \pmod{5} \\ q(B) &= q(U) + 2q(V) \pmod{5} \end{aligned}$$

Construction of design

Column (level of V)

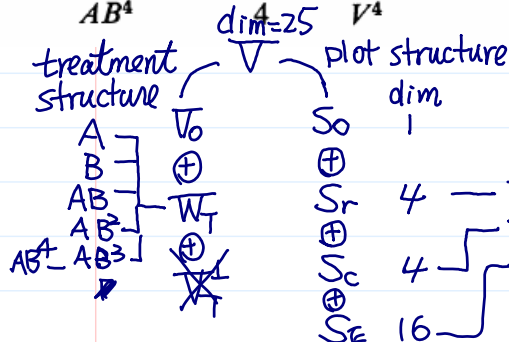
0 1 2 3 4

Levels of A (first) and B (second)

Row (level of U)	0	1	2	3	4
0	00	12	24	31	43
1	11	23	30	42	04
2	22	34	41	03	10
3	33	40	02	14	21
4	44	01	13	20	32

Treatment effect	d.f.	Plot alias	Stratum
A	4	UV	UV
B	4	UV^2	UV
AB	4	$U^2 V^3$	UV
AB^2	4	U^3	U
AB^3	4	$U^4 V^2$	UV
AB^4	4	V^4	V

S_E



$$\begin{aligned} AB &= (UV)(UV^2) = U^2 V^3 \approx UV^4 \\ AB^2 &= UV \cdot U^2 V^4 = U^3 V^5 = U^3 \approx U \\ AB^3 &= \dots \end{aligned}$$

$$AB^4 = \dots$$

6-13

A single replicate block design

$(N=t)$

Treatment structure

$$A \times B \times C \times D$$

full factorial

Design key

$$\begin{array}{cccc} A & B & C & D \\ Y_1 & Y_2 & X_1 Y_1 Y_2 & X_2 Y_1 Y_2 \end{array}$$

Inverse key

$$\begin{array}{cc} X_1 & X_2 \\ ABC & ABD \end{array}$$

Stratum contents

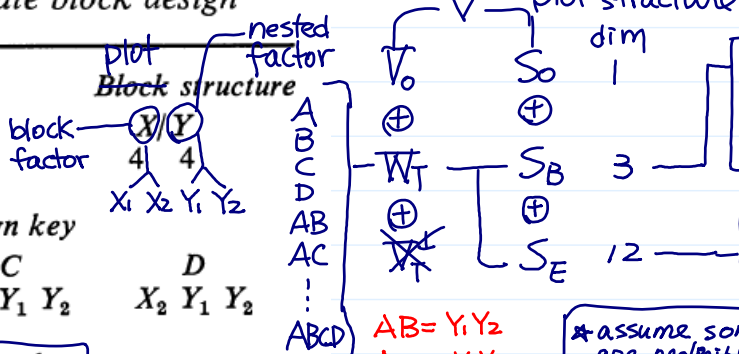
Stratum	d.f.	Treatment effects
X	3	ABC, ABD, CD
XY	12	The rest

Partial inverse key

$$\begin{array}{cc} X_1 & X_2 \\ ABC & ABD \end{array}$$

Rules of construction

$$\begin{aligned} q(A) &= q(Y_1) \pmod{2} \\ q(B) &= q(Y_2) \pmod{2} \\ q(C) &= q(X_1) + q(Y_1) + q(Y_2) \pmod{2} \\ q(D) &= q(X_2) + q(Y_1) + q(Y_2) \pmod{2} \end{aligned}$$



$$AB = Y_1 Y_2$$

$$AC = X_1 Y_2$$

$$ABC = Y_1 Y_2 X_1 Y_1 Y_2 = X_1$$

$$ABD = X_2$$

$$CD = X_1 X_2$$

* assume some higher-order interactions are negligible, some d.f. are left for testing

design matrix

	ABC	ABD	
X_1	0	0	0
X_2	0	0	0
Y_1	0	0	0
Y_2	0	0	0
$X_1 Y_1$	0	0	0
$X_1 Y_2$	0	0	0
$X_2 Y_1$	0	0	0
$X_2 Y_2$	0	0	0
$X_1 Y_1 Y_2$	0	0	0
$X_2 Y_1 Y_2$	0	0	0
$X_1 X_2 Y_1$	0	0	0
$X_1 X_2 Y_2$	0	0	0
$X_1 X_2 Y_1 Y_2$	0	0	0
$X_1 X_2 Y_1 Y_2$	0	0	0

main effects A, B
 confounded with plot main effects

a list of treatment effects whose plot aliases are plot main effects

6-14

A design with crossing and nesting

$(N=t)$

Treatment structure: $A \times B \times C$
3 3 3

Block structure: $(U \times V)/W$
3 3 3

full factorial

Design key: A B C
W UVW $UV^2 W$

Stratum identification

Treatment effect	d.f.	Plot alias
A	2	W
B	2	UVW
C	2	$UV^2 W$
AB	2	UVW^2
AB^2	2	$U^2 V^2$
AC	2	$UV^2 W^2$
AC^2	2	$U^2 V$
BC	2	$U^2 W^2$
BC^2	2	V^2
ABC	2	U^2
ABC^2	2	$V^2 W$
$AB^2 C$	2	VW
$AB^2 C^2$	2	UW^2

Inverse key

U V W
 $A^2 B^2 C^2$ $B^2 C$ A

Partial inverse key

U V
 $A^2 B^2 C^2$ $B^2 C$

$$N(c(n_1, n_2, n_3)) = 1 + (c(n_1, n_2) - 1) + \frac{n_1 n_2}{c(n_1, n_2)} (n_3 - 1)$$

$$= 1 + (n_1 - 1) + (n_2 - 1) + \frac{n_1 n_2}{c(n_1, n_2)} (n_3 - 1)$$

$$= \frac{1}{S_0} + \frac{(n_1 - 1)}{S_1} + \frac{(n_2 - 1)}{S_2} + \frac{(n_1 - 1)(n_2 - 1)}{S_3} + \frac{(n_1 n_2 - 1)}{S_4} (n_3 - 1)$$

plot structure

drn

S_0 1

S_1 2 - U

S_2 2 - V

S_3 4 - UV
 UV^2

S_4 18 - other plot effects with W.

$$AB = W \cdot UVW = UVW^2 \in S_4$$

design matrix

6-15

$(N=t)$

A half-replicate design

full: 32 run
 $\frac{1}{2}$ FFD: 16 run

Treatment structure: $A \times B \times C \times D \times E$
2 2 2 2 2

Block structure: $X(Y)$
2 8

Design key: A B C D E
 $Y_1 Y_2 Y_3 XY_1 Y_2$

Inverse key

X $Y_1 Y_2 Y_3$ I
ABD A B C ABCE

Partial inverse key

X I
ABD ABCE

\rightarrow # of plots = 16

32 all treatment effects.

A = BCE
B = ACE
C =
D =
E =
AB =
AC =

ABCDE

drn

V_0 S_0 1

W_T S_B 1 - X

S_E 14 - $\begin{bmatrix} Y_1 \\ Y_2 \\ Y_3 \\ XY_1 \\ XY_2 \\ XY_3 \\ Y_1 Y_2 \\ Y_1 Y_3 \\ Y_2 Y_3 \end{bmatrix}$

BCE = A = Y_1
ACE = B = Y_2
= C = Y_3
D = $XY_1 Y_2$
AB = $Y_1 Y_2$
AC = --

CDE = ABD = X
ABCD =

fixed effect approach,
only the alias set ABD = CDE are
sacrificed for block effects.

* Split-plot design
the assignment
of design key
is restricted

6-16