

# Normal Distribution and Sequential ANOVA

Consider a linear model  $\underline{Y} = \underline{X}\underline{\beta} + \underline{\varepsilon}$ , where

may have more than one effects

$\underline{Y} \in \mathbb{R}^n$ : an  $n \times 1$  random vector and  $\underline{Y} \sim N(\underline{X}\underline{\beta}, \sigma^2 I)$

Type I ANOVA

$n \times 1$  vectors

sample size

$$\underline{X} = \begin{bmatrix} 1 & \underline{X}_1 & \dots & \underline{X}_k \end{bmatrix}$$

intercept

$$\underline{\beta} = \begin{bmatrix} \beta_0^T & \beta_1^T & \dots & \beta_k^T \end{bmatrix}^T$$

$$\underline{X}\underline{\beta} = \underbrace{\begin{bmatrix} 1 \\ \underline{X}_0 \end{bmatrix}}_{\mu_0} \beta_0 + \underbrace{\underline{X}_1}_{\mu_1} \beta_1 + \dots + \underbrace{\underline{X}_k}_{\mu_k} \beta_k$$

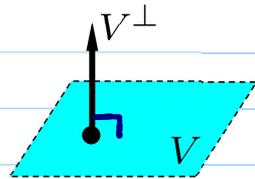
sequential ANOVA

$\underline{\varepsilon} \sim N(\underline{0}, \sigma^2 I) \leftarrow \varepsilon_1, \dots, \varepsilon_n \text{ i.i.d. } N(0, \sigma^2)$

$\mu_i \in \text{span}\{\underline{X}_i\}$

Define  $H_0^{(i)}: \beta_i = 0 \text{ (} \underline{V}_{i-1} \text{) vs. } H_i^{(i)}: \beta_i \neq 0 \text{ (} \underline{V}_i \text{)}, i=1, \dots, k.$

$\underline{V}_i$  = the vector space generated by the column vectors of



$$\underline{A}_i = [\underline{A}_{i-1} \ \underline{X}_i] \quad \underline{A}_i \equiv [\underline{X}_0 \ \underline{X}_1 \ \dots \ \underline{X}_i], \quad i=0, 1, \dots, k.$$

\*  $\underline{V}_i$  is called the *model space* of  $\underline{A}_i$ , and denoted by  $\text{span}\{\underline{A}_i\}$

\*  $V_0 \subset V_1 \subset \dots \subset V_k = \text{span}\{\underline{X}\} \subset \mathbb{R}^n$

$\underline{V}^\perp$  = orthogonal complement of the vector space  $\underline{V}$ , i.e.,  $\mathbb{R}^n = \underline{V} \oplus \underline{V}^\perp$

a vector space

$$\underline{V}^\perp = \{ \underline{v} \in \mathbb{R}^n : \underline{v} \text{ is orthogonal to all the vectors in } \underline{V} \}$$

check graph in LNp.2-25

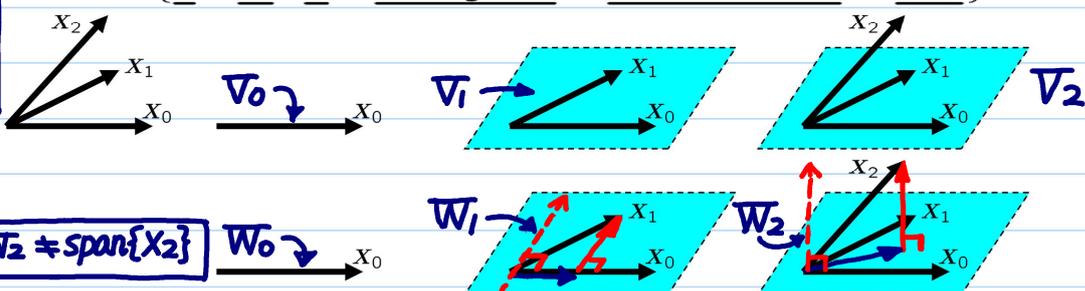
$\underline{W}_0 = \text{span}\{\underline{X}_0\} = \underline{V}_0$ , and for  $i=1, \dots, k$ ,  $\underline{W}_i$  = orthogonal complement of  $\underline{V}_{i-1}$  relative to  $\underline{V}_i$  (note:  $\underline{V}_{i-1} \subset \underline{V}_i$ ), i.e.,

$\underline{W}_i$ : 隨機 ( $H_0^{(i)}$ ) or 規律 ( $H_i^{(i)}$ )? for testing  $H_0^{(i)}: \beta_i = 0$  in type I ANOVA.

a vector space

$$\underline{W}_i = \underline{V}_i \ominus \underline{V}_{i-1} = \underline{V}_i \cap \underline{V}_{i-1}^\perp \iff \underline{V}_i \cap \underline{V}_{i-1}^\perp \neq \underline{W}_i \text{ in general}$$

$\text{span}\{\underline{A}_i\} \ominus \text{span}\{\underline{A}_{i-1}\} = \text{span}\{[\underline{A}_{i-1} \ \underline{X}_i]\} \ominus \text{span}\{\underline{A}_{i-1}\} \neq \text{span}\{\underline{X}_i\}$  in general



$\underline{W}_1 \neq \text{span}\{\underline{X}_1\}, \underline{W}_2 \neq \text{span}\{\underline{X}_2\}$

direct sum

\*  $\underline{V}_i = \underline{W}_0 \oplus \underline{W}_1 \oplus \dots \oplus \underline{W}_i \Rightarrow \underline{v}$  is uniquely represented =  $\underline{w}_0 + \underline{w}_1 + \dots + \underline{w}_i$

\*  $\mathbb{R}^n = \underline{W}_0 \oplus \underline{W}_1 \oplus \dots \oplus \underline{W}_k \oplus \underline{V}_k^\perp$  (note:  $\underline{V}_k^\perp$  is the residual space)

\*  $\underline{W}_0 \perp \underline{W}_1 \perp \dots \perp \underline{W}_k \perp \underline{V}_k^\perp \Rightarrow$  orthogonality

mutually orthogonal under the linear model in LNp.31.

usually equals the # of effects in  $\underline{X}_i$

\* Note. In general,  $\underline{W}_i \neq \text{span}\{\underline{X}_i\}$ . However, if  $\underline{X}_0, \underline{X}_1, \dots, \underline{X}_k$  are mutually orthogonal, then  $\underline{W}_i = \text{span}\{\underline{X}_i\}$ .

$\underline{r}_i \equiv \dim(\underline{W}_i) = \dim(\underline{V}_i) - \dim(\underline{V}_{i-1})$

\* Let  $\underline{r} = \sum_{i=0}^k \underline{r}_i = \sum_{i=0}^k \dim(\underline{W}_i) = \dim(\underline{V}_k)$ . Then,  $\dim(\underline{V}_k^\perp) = n - \underline{r}$

• Orthogonal projection of  $\underline{Y}$  onto  $V_i$ 's and  $W_i$ 's ← a linear transformation of  $\underline{Y}$  on an  $n \times n$  matrix

For a vector space  $V \subset \mathbb{R}^n$ , denote the orthogonal projection matrix of  $\underline{Y}$  onto  $V$  by  $\underline{P}_V$ . Then, the orthogonal projection of  $\underline{Y}$  onto  $V$  is  $\underline{P}_V \underline{Y}$ .

\* if  $V = \text{span}\{\underline{A}\}$ , then  $\underline{P}_V = \underline{A}(\underline{A}^T \underline{A})^{-1} \underline{A}^T$  ← Recall. Hat matrix in linear model

\* the orthogonal projection matrix onto  $V^\perp$ , denoted by  $\underline{P}_{V^\perp}$ , is  $\underline{P}_{V^\perp} = \underline{I} - \underline{A}(\underline{A}^T \underline{A})^{-1} \underline{A}^T = \underline{I} - \underline{P}_V$

Some properties of orthogonal projection matrix

- A square matrix  $\underline{P}$  is a projection matrix iff  $\underline{P}^2 = \underline{P}$  (idempotent) iff
  - Idempotence implies  $\underline{P}$  is a generalized inverse of  $\underline{P}$  since  $\underline{P} \underline{P} \underline{P} = \underline{P}^3 = \underline{P}$ . (usually, a singular matrix)
- A projection matrix  $\underline{P}$  is orthogonal iff  $\underline{P}^T = \underline{P}$  (symmetric)
- If  $\underline{P}$  is an orthogonal projection matrix onto  $V$ , then
  - $\underline{P}$  has  $\text{dim}(V)$  eigenvalues equal to 1 and the rest 0
  - $\underline{P}$  is diagonalizable, and there exists an orthogonal matrix  $\underline{U}$  ( $\underline{U}^T \underline{U} = \underline{I}$ ) such that  $\underline{U}^T \underline{P} \underline{U} = \underline{\Lambda}$  is a diagonal matrix.  $\underline{\Lambda} = \begin{bmatrix} \lambda_1 & & 0 \\ & \ddots & \\ 0 & & \lambda_n \end{bmatrix}$ ,  $\lambda_i \in \{0, 1\}$
  - (Note. Thus,  $\underline{P} = \underline{U} \underline{\Lambda} \underline{U}^T$ ) Actually,
    - the columns of  $\underline{U}$  are orthonormal eigenvectors of  $\underline{P}$ , and
    - the diagonal entries of  $\underline{\Lambda}$  are the eigenvalues of  $\underline{P}$

rank( $\underline{P}$ ) =  $\text{dim}(V)$

eigen-decomposition

Since  $V_i = \text{span}\{\underline{A}_i\}$ ,  $\underline{P}_{V_i} = \underline{A}_i(\underline{A}_i^T \underline{A}_i)^{-1} \underline{A}_i^T$  and  $\underline{P}_{V_i} \underline{Y} = \underline{A}_i(\underline{A}_i^T \underline{A}_i)^{-1} \underline{A}_i^T \underline{Y}$

\* For the model space  $V_i$ , ← i.e. consider the model:  $\underline{Y} \sim \underline{x}_0 \beta_0 + \underline{x}_1 \beta_1 + \dots + \underline{x}_i \beta_i$

$\text{RSS}_{V_i} = \|\underline{Y}\|^2 - \|\underline{P}_{V_i} \underline{Y}\|^2 = \underline{Y}^T \underline{Y} - (\underline{P}_{V_i} \underline{Y})^T \underline{P}_{V_i} \underline{Y} = \underline{Y}^T \underline{Y} - \underline{Y}^T \underline{P}_{V_i}^T \underline{P}_{V_i} \underline{Y}$

$= \underline{Y}^T \underline{Y} - \underline{Y}^T \underline{P}_{V_i} \underline{Y} = \underline{Y}^T (\underline{I} - \underline{P}_{V_i}) \underline{Y} = \|(I - P_{V_i})Y\|^2 = \underline{P}_{V_i}^T \underline{P}_{V_i} = \underline{P}_{V_i}$

Since  $W_i = V_i \cap V_{i-1}^\perp$  and  $V_{i-1} \subset V_i$ ,  $\underline{P}_{W_i} = (\underline{I} - \underline{P}_{V_{i-1}}) \underline{P}_{V_i} = \underline{P}_{V_i} - \underline{P}_{V_{i-1}} \underline{P}_{V_i} = \underline{P}_{V_i} - \underline{P}_{V_{i-1}}$

and  $\underline{P}_{W_i} \underline{Y} = (\underline{I} - \underline{P}_{V_{i-1}}) \underline{P}_{V_i} \underline{Y} = (\underline{P}_{V_i} - \underline{P}_{V_{i-1}}) \underline{Y} = \underline{P}_{V_i} \underline{Y} - \underline{P}_{V_{i-1}} \underline{Y}$

\* Since  $\mathbb{R}^n = W_0 \oplus W_1 \oplus \dots \oplus W_k \oplus V_k^\perp$  and  $W_0 \perp W_1 \perp \dots \perp W_k \perp V_k^\perp$ ,

$\underline{Y} = \underline{P}_{W_0} \underline{Y} + \underline{P}_{W_1} \underline{Y} + \dots + \underline{P}_{W_k} \underline{Y} + \underline{P}_{V_k^\perp} \underline{Y}$

$\underline{P}_{W_0} \underline{Y} \perp \underline{P}_{W_1} \underline{Y} \perp \dots \perp \underline{P}_{W_k} \underline{Y} \perp \underline{P}_{V_k^\perp} \underline{Y}$

$\|\underline{Y}\|^2 = \|\underline{P}_{W_0} \underline{Y}\|^2 + \|\underline{P}_{W_1} \underline{Y}\|^2 + \dots + \|\underline{P}_{W_k} \underline{Y}\|^2 + \|\underline{P}_{V_k^\perp} \underline{Y}\|^2$

\* When  $\underline{X}_0, \underline{X}_1, \dots, \underline{X}_k$  are mutually orthogonal,  $\underline{W}_i = \text{span}\{\underline{X}_i\}$

$\underline{A}_i(\underline{A}_i^T \underline{A}_i)^{-1} \underline{A}_i^T - \underline{A}_{i-1}(\underline{A}_{i-1}^T \underline{A}_{i-1})^{-1} \underline{A}_{i-1}^T = \underline{P}_{W_i} = \underline{P}_{V_i} - \underline{P}_{V_{i-1}} = \underline{X}_i(\underline{X}_i^T \underline{X}_i)^{-1} \underline{X}_i^T$   $\underline{A}_i = [\underline{A}_{i-1} \ \underline{X}_i]$

check the graphs in LNp.32

check  $\star$  in LNp.32

Partition of sums of squares in Type I ANOVA

Type III ANOVA