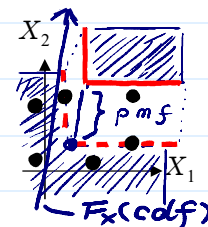


(iv) $F_{\mathbf{X}}$ is continuous from the right with respect to each of the coordinates, or any subset of them jointly, i.e., if $\mathbf{x} = (x_1, \dots, x_n)$ and $\mathbf{z}_m = (z_{1m}, \dots, z_{nm})$ such that $\mathbf{z}_m \downarrow \mathbf{x}$, then



$$F_{\mathbf{X}}(\mathbf{z}_m) \downarrow F_{\mathbf{X}}(\mathbf{x}) \quad (\text{Ex})$$

$$\begin{aligned} & \bullet (-\infty, \mathbf{z}_m] \\ & \equiv (-\infty, z_{1m}] \times \dots \times (-\infty, z_{nm}] \\ & \bullet (-\infty, \mathbf{z}_m] \downarrow \\ & \quad (-\infty, \mathbf{x}] \end{aligned}$$

$$\begin{aligned} & (-\infty, x_1] \times \dots \times (-\infty, x_n] \\ & \subseteq (-\infty, t_1] \times \dots \times (-\infty, t_n] \\ & \subseteq (-\infty, x'_1] \times \dots \times (-\infty, x'_n] \end{aligned} \quad (\text{Ex})$$

non-decreasing (cf., (2) in LNp.5-9)

$$F_{\mathbf{X}}(x_1, \dots, x_n) \leq F_{\mathbf{X}}(t_1, \dots, t_n) \leq F_{\mathbf{X}}(x'_1, \dots, x'_n) \quad (\text{Ex})$$

where $t_i \in \{x_i, x'_i\}$, $i = 1, 2, \dots, n$. When $n=2$, we have

$$F_{X_1, X_2}(x_1, x_2) \leq \begin{cases} F_{X_1, X_2}(x_1, x'_2) \\ F_{X_1, X_2}(x'_1, x_2) \end{cases} \leq F_{X_1, X_2}(x'_1, x'_2).$$

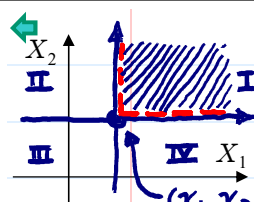
Q how to generalize to any n ? cf. (5) in LNp.5-10

(vi) If $x_1 \leq x'_1$ and $x_2 \leq x'_2$, then

inclusion-exclusion identity (LNp.3-12)

$$\text{Blue} \cap (\text{Red} \cup \text{Green})^c = \text{Blue} \setminus (\text{Red} \cup \text{Green})$$

$$\begin{aligned} & (\Delta) - P(x_1 \leq X_1 \leq x'_1, x_2 \leq X_2 \leq x'_2) \\ & = F_{X_1, X_2}(x'_1, x'_2) - F_{X_1, X_2}(x_1, x'_2) \\ & \quad - F_{X_1, X_2}(x'_1, x_2) + F_{X_1, X_2}(x_1, x_2) \end{aligned}$$



In particular, let $x'_1 \uparrow \infty$ and $x'_2 \uparrow \infty$, we get

$$P(x_1 < X_1 < \infty, x_2 < X_2 < \infty) = 1 - F_{X_1}(x_1) - F_{X_2}(x_2) + F_{X_1, X_2}(x_1, x_2).$$

can be any R r.v.'s

(vii) The joint cdf of $X_1, \dots, X_k$, $k < n$, is

$$F_{X_1, \dots, X_k}(x_1, \dots, x_k) = P(X_1 \leq x_1, \dots, X_k \leq x_k)$$

$$= P(X_1 \leq x_1, \dots, X_k \leq x_k, -\infty < X_{k+1} < \infty, \dots, -\infty < X_n < \infty)$$

$$\lim_{x_{k+1}, x_{k+2}, \dots, x_n \rightarrow \infty} F_{\mathbf{X}}(x_1, \dots, x_k, x_{k+1}, \dots, x_n).$$

In particular, the marginal cdf of X_1 is

Sometimes, might call it marginal joint cdf or marginal joint distribution (abusing the terminologies)

$$F_{X_1}(x) = P(X_1 \leq x) = \lim_{x_2, x_3, \dots, x_n \rightarrow \infty} F_{\mathbf{X}}(x, x_2, x_3, \dots, x_n).$$

Theorem in LNp.5-11

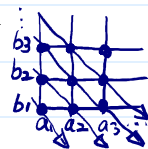
Theorem. A function $F_{\mathbf{X}}(x_1, \dots, x_n)$ can be a joint cdf if $F_{\mathbf{X}}$ satisfies (i)-(v) in the previous theorem.

Joint Probability Mass Function

■ Definition. Suppose that $X_1, \dots, X_n$ are discrete random variables. The joint pmf of $\mathbf{X} = (X_1, \dots, X_n)$ is defined as

$$p_{\mathbf{X}}: \mathbb{R}^n \rightarrow \mathbb{R}$$

$$p_{\mathbf{X}}(x_1, \dots, x_n) = P(X_1 = x_1, \dots, X_n = x_n).$$



■ Theorem. Suppose that $p_{\mathbf{X}}$ is a joint pmf. Then,

(a) $p_{\mathbf{X}}(x_1, \dots, x_n) \geq 0$, for $-\infty < x_i < \infty$, $i = 1, \dots, n$.

(b) There exists a finite or countably infinite set $\mathcal{X} \subset \mathbb{R}^n$ such that $p_{\mathbf{X}}(x_1, \dots, x_n) = 0$, for $(x_1, \dots, x_n) \notin \mathcal{X}$.

(c) $\sum_{\mathbf{x} \in \mathcal{X}} p_{\mathbf{X}}(\mathbf{x}) = 1$, where $\mathbf{x} = (x_1, \dots, x_n)$.

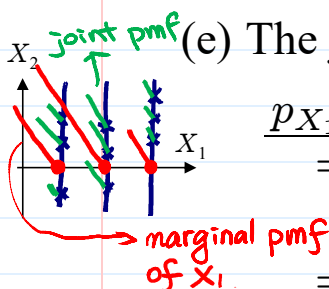
(d) For $A \subset \mathbb{R}^n$, $P(\mathbf{X} \in A) = \sum_{\mathbf{x} \in A \cap \mathcal{X}} p_{\mathbf{X}}(\mathbf{x})$.

(e) The joint pmf of $X_1, \dots, X_k$, $k < n$, is

$$p_{X_1, \dots, X_k}(x_1, \dots, x_k) = P(X_1 = x_1, \dots, X_k = x_k)$$

$$= P(X_1 = x_1, \dots, X_k = x_k, \underbrace{-\infty < X_{k+1} < \infty, \dots, -\infty < X_n < \infty}_{\text{same event}})$$

$$= \sum_{\substack{(x_1, \dots, x_n) \in \mathcal{X} \\ -\infty < x_{k+1} < \infty, \dots, -\infty < x_n < \infty}} p_{\mathbf{X}}(x_1, \dots, x_k, x_{k+1}, \dots, x_n).$$



joint $\Rightarrow$ marginal

$$\sum_{\substack{(x_1, \dots, x_n) \in \mathcal{X} \\ -\infty < x_{k+1} < \infty, \dots, -\infty < x_n < \infty}}$$

In particular, the marginal pmf of X_1 is

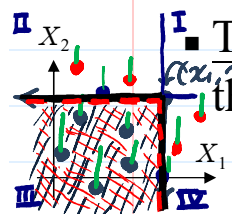
check the example in LNp.7-3

$$p_{X_1}(x) = P(X_1 = x)$$

$$= \sum_{\substack{(x, x_2, \dots, x_n) \in \mathcal{X} \\ -\infty < x_2 < \infty, \dots, -\infty < x_n < \infty}} p_{\mathbf{X}}(x, x_2, x_3, \dots, x_n).$$

can be any one of the n r.v.'s

■ Theorem. A function $p_{\mathbf{X}}(x_1, \dots, x_n)$ can be a joint pmf if $p_{\mathbf{X}}$ satisfies (a)-(c) in the previous theorem.



■ Theorem. If $F_{\mathbf{X}}$ and $p_{\mathbf{X}}$ are the joint cdf and joint pmf of $\mathbf{X}$, then

$$F_{\mathbf{X}}(x_1, \dots, x_n) = \sum_{\substack{(t_1, \dots, t_n) \in \mathcal{X} \\ t_1 \leq x_1, \dots, t_n \leq x_n}} p_{\mathbf{X}}(t_1, \dots, t_n).$$

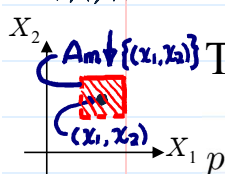
$$\stackrel{P(\mathbf{X} \in A)}{=} \sum_{\substack{(t_1, \dots, t_n) \in \mathcal{X} \\ t_1 \leq x_1, \dots, t_n \leq x_n}} p_{\mathbf{X}}(t_1, \dots, t_n)$$

$$A = (-\infty, x_1] \times (-\infty, x_2] \times \dots \times (-\infty, x_n]$$

$$P_{\mathbf{X}}(x_1, \dots, x_n) \neq P(X_1 \leq x_1, \dots, X_n \leq x_n)$$

$$= P(X_1 < x_1, \dots, X_n < x_n)$$

Recall: $P(x) = F(x) - F(x-)$ in LNp.5-11

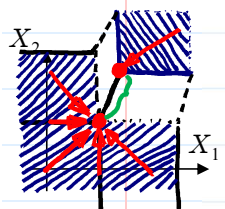


To derive $p_{\mathbf{X}}$ from $F_{\mathbf{X}}$, take $n=2$ to illustrate:

$$p_{\mathbf{X}}(x_1, x_2) = \lim_{m \rightarrow \infty} P\left(x_1 - \frac{1}{m} < X_1 \leq x_1 + \frac{1}{m}, x_2 - \frac{1}{m} < X_2 \leq x_2 + \frac{1}{m}\right)$$

$$= \lim_{m \rightarrow \infty} \left[F_{\mathbf{X}}\left(x_1 + \frac{1}{m}, x_2 + \frac{1}{m}\right) - F_{\mathbf{X}}\left(x_1 + \frac{1}{m}, x_2 - \frac{1}{m}\right) - F_{\mathbf{X}}\left(x_1 - \frac{1}{m}, x_2 + \frac{1}{m}\right) + F_{\mathbf{X}}\left(x_1 - \frac{1}{m}, x_2 - \frac{1}{m}\right) \right]$$

$$= F_{\mathbf{X}}(x_1, x_2) - F_{\mathbf{X}}(x_1, x_2-) - F_{\mathbf{X}}(x_1-, x_2) + F_{\mathbf{X}}(x_1-, x_2-)$$



Note that
pmf $\leftrightarrow$ pdf
 $\Sigma \leftrightarrow \int$

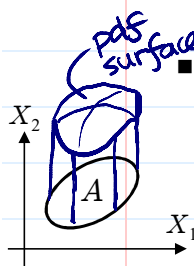
Joint Probability Density Function $\mathbf{X}: \Omega \rightarrow \mathbb{R}^n, f_{\mathbf{X}}: \mathbb{R}^n \rightarrow \mathbb{R}$

■ Definition. A function $f_{\mathbf{X}}(x_1, \dots, x_n)$ can be a joint pdf if

(1) $f_{\mathbf{X}}(x_1, \dots, x_n) \geq 0$, for $-\infty < x_i < \infty, i=1, \dots, n$.

(2) $\int_{-\infty}^{\infty} \cdots \int_{-\infty}^{\infty} f_{\mathbf{X}}(x_1, \dots, x_n) dx_1 \cdots dx_n = 1$.

If $\dim(A) < n$
 $P(\mathbf{X} \in A) = 0$



■ Definition. Suppose that $X_1, \dots, X_n$ are continuous r.v.'s.

The joint pdf of $\mathbf{X} = (X_1, \dots, X_n)$ is a function $f_{\mathbf{X}}(x_1, \dots, x_n)$ satisfying (1) and (2) above, and for any event $A \subset \mathbb{R}^n$,

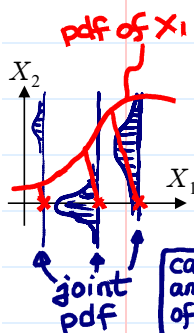
$P(\mathbf{X} \in A) = \int \cdots \int_A f_{\mathbf{X}}(x_1, \dots, x_n) dx_1 \cdots dx_n$.
In practice, could be difficult to identify A for the event of interest

joint $\Rightarrow$ marginal

■ Theorem. Suppose that $f_{\mathbf{X}}$ is the joint pdf of $\mathbf{X} = (X_1, \dots, X_n)$.

Then, the joint pdf of $X_1, \dots, X_k, k < n$, is

can be any R.v.'s in $\mathbf{X}$



$f_{X_1, \dots, X_k}(x_1, \dots, x_k) = \int_{-\infty}^{\infty} \cdots \int_{-\infty}^{\infty} f_{\mathbf{X}}(x_1, \dots, x_k, x_{k+1}, \dots, x_n) dx_{k+1} \cdots dx_n$.
by (*) in LNP 7-11 $\frac{\partial^k}{\partial x_1 \cdots \partial x_k} F_{\mathbf{X}}(x_1, \dots, x_n) = \frac{\partial^k}{\partial x_1 \cdots \partial x_k} \left(\lim_{x_{k+1} \rightarrow \infty} \cdots \lim_{x_n \rightarrow \infty} F_{X_1, \dots, X_k, X_{k+1}, \dots, X_n} \right)$
by (o) in LNP 7-7 $\rightarrow \infty$

In particular, the marginal pdf of X_1 is

$\int_{-\infty}^{\infty} \cdots \int_{-\infty}^{\infty} f_{\mathbf{X}}(x_1, \dots, x_n) dx_2 \cdots dx_n$

can be any one of the n r.v.'s

$f_{X_1}(x) = \int_{-\infty}^{\infty} \cdots \int_{-\infty}^{\infty} f_{\mathbf{X}}(x, x_2, \dots, x_n) dx_2 \cdots dx_n$.

■ Theorem. If $F_{\mathbf{X}}$ and $f_{\mathbf{X}}$ are the joint cdf and joint pdf of $\mathbf{X}$,

then $P(\mathbf{X} \in A) = \int \cdots \int_A f_{\mathbf{X}}(t_1, \dots, t_n) dt_1 \cdots dt_n$

continuous

(*) $\left\{ \begin{aligned} F_{\mathbf{X}}(x_1, \dots, x_n) &= \int_{-\infty}^{x_1} \cdots \int_{-\infty}^{x_n} f_{\mathbf{X}}(t_1, \dots, t_n) dt_1 \cdots dt_n, \text{ and} \\ f_{\mathbf{X}}(x_1, \dots, x_n) &= \frac{\partial^n}{\partial x_1 \cdots \partial x_n} F_{\mathbf{X}}(x_1, \dots, x_n). \end{aligned} \right.$

at the continuity points of $f_{\mathbf{X}}$.

Examples.

➤ Experiment. Two balls are drawn without replacement from a

box with one ball labeled 1,

6 balls

two balls labeled 2,

three balls labeled 3.

① ② ② ③ ③ ③
 $b_1 b_2 b_3 b_4 b_5 b_6$

cf. black & white balls

$\Omega = \{(b_1, b_2), (b_1, b_3), \dots, (b_6, b_5)\}$
 $\#\Omega = 6 \times 5 = 30$, symmetric outcome

$b_1 \dots b_6$
1st 2nd

$\mathbf{X}: \Omega \rightarrow \mathbb{R}$
 $\mathbf{Y}: \Omega \rightarrow \mathbb{R}$

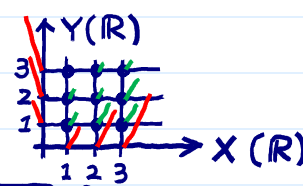
Let

$X =$ label on the 1st ball drawn,

$Y =$ label on the 2nd ball drawn.

jointly distributed r.v.'s

■ The joint pmf and marginal pmfs of (X, Y) are



discrete

joint pmf $\rightarrow$ $p(x, y)$

		X			$p_Y(y)$
		1	2	3	
Y	1	0	2/30	3/30	1/6
	2	2/30	2/30	6/30	2/6
	3	3/30	6/30	6/30	3/6
		$p_X(x)$			
		1/6	2/6	3/6	

marginal pmf of X $\rightarrow$ $p_X(x)$

marginal pmf of Y $\rightarrow$ $p_Y(y)$

$P(X=3, Y=3)$
 $= P(X=3)P(Y=3|X=3)$
 $= \frac{3}{6} \times \frac{2}{5} = \frac{6}{30} \leftarrow \# \Omega$

Q: The balls are drawn without replacement. Why do X (from 1st ball) and Y (from 2nd ball) have same marginal distributions?

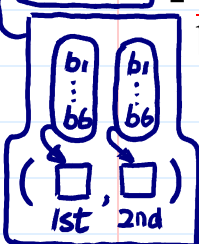
$Z = |X - Y|$
 $Z: \Omega \rightarrow \mathbb{R}$
 $P(Z=1) = ?$

Q: $P(|X - Y| = 1) = ??$ need joint dist

transformation of (X, Y)

$$P(|X - Y| = 1) = P(X=1, Y=2) + P(X=2, Y=1) + P(X=2, Y=3) + P(X=3, Y=2) = 8/15.$$

$\# \Omega = 36$



Q: What are the joint pmf and marginal pmfs of (X, Y) if the balls are drawn with replacement (LNp. 4-6)?

$$P(|X - Y| = 1) = \frac{4}{9}$$

		X			$p_Y(y)$
		1	2	3	
Y	1	1/36	2/36	3/36	1/6
	2	2/36	4/36	6/36	2/6
	3	3/36	6/36	9/36	3/6
		$p_X(x)$			
		1/6	2/6	3/6	

Same marginal pmf as without replacement

$$p(x, y) = p_X(x)p_Y(y)$$

$\rightarrow$ independent

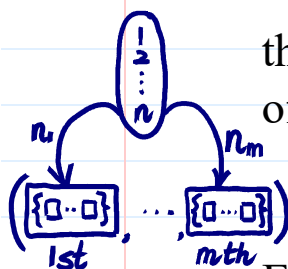
➤ Multinomial Distribution $\xleftrightarrow{\text{c.f.}}$ binomial distribution

■ Recall. Partitions

□ If $n \geq 1$ and $n_1, \dots, n_m \geq 0$ are integers for which

$$n_1 + \dots + n_m = n,$$

then a set of n elements may be partitioned into m subsets of sizes $n_1, \dots, n_m$ in

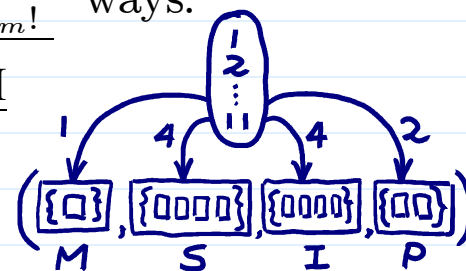


$$\binom{n}{n_1, \dots, n_m} = \frac{n!}{n_1! \times \dots \times n_m!} \text{ ways.}$$

□ Example (LNp.2-8) : MISSISSIPPI

$$\binom{11}{4, 1, 2, 4} = \frac{11!}{4!1!2!4!}.$$

■ Example (Die Rolling) $\xleftrightarrow{\text{c.f.}}$ Coin Tossing



□ **Q:** If a balanced (6-sided) die is rolled 12 times,

各种結果出現次數

$\rightarrow P(\text{each face appears twice}) = ??$

相同實驗, 不斷重複

□ Sample space of rolling the die once (basic experiment):

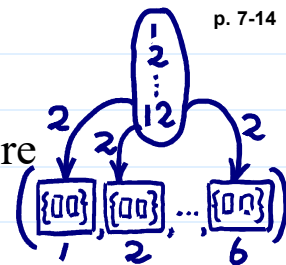
$$\Omega_0 = \{1, 2, 3, 4, 5, 6\}. \leftarrow \text{symmetric outcomes}$$

- The sample space for the 12 trials is

$$\underline{\Omega} = \Omega_0 \times \cdots \times \Omega_0 = \underline{\Omega}_0^{12}$$

An outcome $\underline{\omega} \in \underline{\Omega}$ is $\underline{\omega} = (i_1, i_2, \dots, i_{12})$, where
 $1 \leq i_1, \dots, i_{12} \leq 6$.

↑ symmetric outcomes
 ∴ balanced die



- There are 6¹² possible outcomes in $\underline{\Omega}$, i.e., $\#\Omega = 6^{12}$.

- Among all possible outcomes, there are $\binom{12}{2,2,2,2,2,2} = \frac{12!}{(2!)^6}$ of which each face appears twice.

□ $P(\text{each face appears twice}) = \frac{12!}{(2!)^6} \left(\frac{1}{6}\right)^{12} = \frac{1}{\#\Omega}$

In $\Omega_0, P(\{1\}) = \dots = P(\{6\}) = 1/6$

■ Generalization.

- Consider a basic experiment which can result in one of m types of outcomes. Denote its sample space as

$$\underline{\Omega}_0 = \{1, 2, \dots, m\} \leftarrow \text{may not be symmetric}$$

Let $p_i = P(\text{outcome } i \text{ appears in a basic experiment})$,

then, (i) $p_1, \dots, p_m \geq 0$, and

(ii) $p_1 + \dots + p_m = 1$.

- Repeat the basic experiment n times. Then, the sample space for the n trials is

$$\underline{\Omega} = \Omega_0 \times \cdots \times \Omega_0 = \underline{\Omega}_0^n \rightarrow \underline{\omega} \in \underline{\Omega}, \underline{\omega}: n\text{-dim vector}$$

discrete
r.v.

Let $X_i = \#$ of trials with outcome i , $i=1, \dots, m$,

Then,

(i) $X_1, \dots, X_m: \underline{\Omega} \rightarrow \mathbb{R}$, and $\underline{X} = (X_1, \dots, X_m): \underline{\Omega} \rightarrow \mathbb{R}^m$

(ii) $X_1 + \dots + X_m = n$. jointly distributed

- The joint pmf of $X_1, \dots, X_m$ is

$$p_{\underline{X}}(x_1, \dots, x_m) = P(X_1 = x_1, \dots, X_m = x_m) = \binom{n}{x_1, \dots, x_m} p_1^{x_1} \times \cdots \times p_m^{x_m}.$$

for $x_1, \dots, x_m \geq 0$ and $x_1 + \dots + x_m = n$.

range $\mathcal{X}$

Proof. The probability of any sequence with x_i 's is

$$p_1^{x_1} \times \cdots \times p_m^{x_m},$$

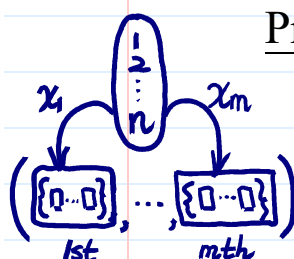
and there are

$$\binom{n}{x_1, \dots, x_m}$$

such sequences. $\leftarrow$ they are symmetric outcomes.

$$\text{Q: } \sum_{\underline{x} \in \mathcal{X}} P_{\underline{X}}(\underline{x}) \neq 1$$

Note. Not saying all $\underline{\omega}$'s in $\underline{\Omega}$ are symmetric

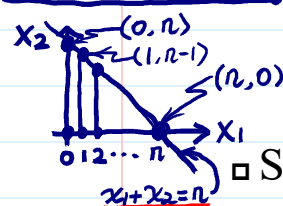


- The distribution of a random vector $\underline{X}=(X_1, \dots, X_m)$ with the above joint pmf is called the multinomial distribution with parameters n, m , and $p_1, \dots, p_m$, denoted by Multinomial $(n, m, p_1, \dots, p_m)$.

- ◆ The multinomial distribution is called after the Multinomial Theorem:

$$\sum_{\underline{x} \in \mathcal{X}} P_{\underline{X}}(\underline{x}) = (p_1 + \dots + p_m)^n = 1$$

$$(a_1 + \dots + a_m)^n = \sum_{\substack{\underline{x} \in \mathcal{X} \\ x_i \in \{0, \dots, n\}; i=1, \dots, m \\ x_1 + \dots + x_m = n}} \binom{n}{x_1, \dots, x_m} a_1^{x_1} \times \dots \times a_m^{x_m} (*)$$

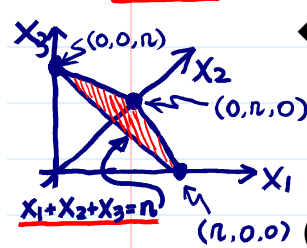
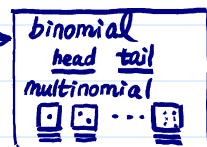


- ◆ It is a generalization of the binomial distribution from 2 types of outcomes to m types of outcomes.

- Some Properties.

$$m=2, (X_1, X_2)$$

$$\text{binomial \& } X_2 = n - X_1$$



- ◆ Because $X_i = n - (X_1 + \dots + X_{i-1} + X_{i+1} + \dots + X_m)$, and $p_i = 1 - (p_1 + \dots + p_{i-1} + p_{i+1} + \dots + p_m)$,

WLOG, we can write $\underline{X} \in \mathbb{R}^m$, but the dimension of its $\mathcal{X} = m-1$

$$(X_1, \dots, X_{m-1}, X_m) \rightarrow (X_1, \dots, X_{m-1}, n - (X_1 + \dots + X_{m-1}))$$