

So, $N_2 \sim \text{Poisson}(11)$, $P(N_2 = k) = \frac{11^k \cdot e^{-11}}{k!}$ and
 2 months \rightarrow

$$P(N_2 \geq 3) = 1 - P(N_2 \leq 2) = 1 - \sum_{k=0}^2 \frac{e^{-11} \cdot 11^k}{k!}$$

Summary for $X \sim \text{Poisson}(\lambda)$

- Range: $\mathcal{X} = \{0, 1, 2, \dots\}$
- Pmf: $f_X(x) = \lambda^x e^{-\lambda} / x!$, for $x \in \mathcal{X}$
- Parameter: $0 < \lambda < \infty$
- Mean: $E(X) = \lambda$
- Variance: $Var(X) = \lambda$

Q: What's the original sample space Ω for the X ?
 For $\omega \in \Omega$, $N_t(\omega)$: a nondecreasing step function of t .

cf. the Ω for binomial X (LNp.5-20~21)
 the Ω for negative binomial X (LNp.5-26~27)

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Hypergeometric Distribution

Experiment: Draw a sample of n ($\leq N$) balls without replacement from a box containing R red balls and $N-R$ white balls (the balls are equally likely to be drawn)

- Let X be the number of red balls in the sample
- Q: What is $P(X=k)$? \leftarrow distribution of X is ?
- Example. The Committee Example (LNp.5-6).
- (cf.) If drawn with replacement, what is the distribution of X ?

binomial($n, \frac{R}{N}$) (check LNp.5-21)

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Recall. LNp.4-5 ~4-7

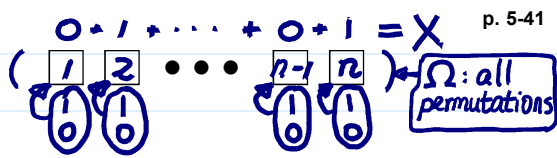
graphs in LNp.5-21

Probability Mass Function

Theorem. For $k = 0, 1, 2, \dots, n$,

$$P(X = k) = \frac{\binom{R}{k} \binom{N-R}{n-k}}{\binom{N}{n}}$$

possible values of X \rightarrow
 intuition \rightarrow



with replacement \leftarrow indep. Bernoulli $\rightarrow X \sim \text{binomial}$
 without replacement \leftarrow dep. Bernoulli (Δ) $\rightarrow X \sim \text{hypergeometric}$ (cf.)

original sample space Ω of X : $\Omega = \{\text{combination of } n\}$
 $\# \Omega = \binom{N}{n}$

(Notice that $\binom{r}{t} \equiv 0$ when either $t < 0$ or $r < t$.)

graphs in LNp.5-20~21

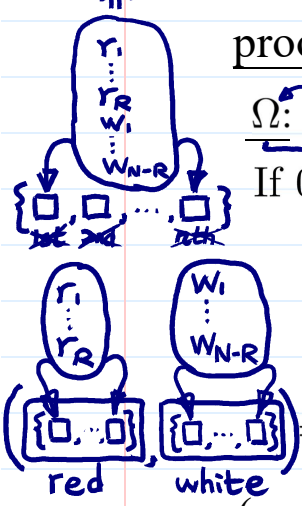
proof. Label the N balls as $r_1, \dots, r_R, w_1, \dots, w_{N-R}$.

Ω : combinations of size n from N different balls. $\Rightarrow \# \Omega = \binom{N}{n}$
 alternative: all $\binom{N}{n}$ permutations.
 symmetric outcomes

If $0 \leq k \leq R$ and $0 \leq n - k \leq N - R$,

k red balls may be chosen in $\binom{R}{k}$ ways.

$n - k$ white balls may be chosen in $\binom{N-R}{n-k}$ ways.



$$\Rightarrow \#\{X = k\} = \binom{R}{k} \binom{N-R}{n-k}$$

(exercise) Show that the following function is a pmf.

For (iii) in LNp.5-6, apply Δ in LNp.5-42

$$f(k) = \begin{cases} \binom{R}{k} \binom{N-R}{n-k} / \binom{N}{n}, & k = 0, 1, \dots, n, \\ 0, & \text{otherwise.} \end{cases}$$

- The distribution of the r.v. X is called the hypergeometric distribution with parameters $n, N,$ and R .

- The hypergeometric distribution is called after the hypergeometric identity:

$$i+j=r$$

$$(1+x)^a(1+x)^b = \left[\sum_{i=0}^a \binom{a}{i} x^i \right] \left[\sum_{j=0}^b \binom{b}{j} x^j \right]$$

$$(1+x)^{a+b} = \sum_{r=0}^{a+b} \binom{a+b}{r} x^r$$

(why? \because (Δ) in LNp.5-41) $\rightarrow \frac{1}{V}$

$$\triangle 3 - \binom{a+b}{r} = \sum_{k=0}^r \binom{a}{k} \binom{b}{r-k}$$

\triangleright Theorem. The mean and variance of hypergeometric(n, N, R) are

intuitive interpretation

$$\mu = \frac{nR}{N} \quad \text{and} \quad \sigma^2 = \frac{nR(N-R)(N-n)}{N^2(N-1)} = n \left(\frac{R}{N} \right) \left(1 - \frac{R}{N} \right) \left(\frac{N-n}{N-1} \right)$$

proof.

same as with replacement (LNp.5-24) \rightarrow with replacement, $\sigma^2 = n \left(\frac{R}{N} \right) \left(1 - \frac{R}{N} \right)$

$$E(X) = \sum_{x=0}^n x \cdot \frac{\binom{R}{x} \binom{N-R}{n-x}}{\binom{N}{n}} = \sum_{x=1}^n x \cdot \frac{R!}{R!} \cdot \frac{N-1}{n-1} \cdot \frac{\binom{N-R}{n-x}}{\binom{N}{n}}$$

$$= \frac{nR}{N} \sum_{x=1}^n \frac{\binom{R-1}{x-1} \binom{(N-1)-(R-1)}{(n-1)-(x-1)}}{\binom{N-1}{n-1}} = \frac{nR}{N} \sum_{y=0}^{n-1} \frac{\binom{R-1}{y} \binom{(N-1)-(R-1)}{(n-1)-y}}{\binom{N-1}{n-1}} = \frac{nR}{N}$$

Let $y=x-1$

pmf of hypergeometric($n-1, N-1, R-1$)

$$E[X(X-1)] = E(X^2 - X) = E(X^2) - E(X)$$

$$= \sum_{x=0}^n x(x-1) \cdot \frac{\binom{R}{x} \binom{N-R}{n-x}}{\binom{N}{n}} = \sum_{x=2}^n x(x-1) \cdot \frac{R!}{R!} \cdot \frac{N-2}{n-2} \cdot \frac{\binom{N-R}{n-x}}{\binom{N}{n}}$$

$$= \frac{n(n-1)R(R-1)}{N(N-1)} \sum_{x=2}^n \frac{\binom{R-2}{x-2} \binom{(N-2)-(R-2)}{(n-2)-(x-2)}}{\binom{N-2}{n-2}}$$

Let $y=x-2$

$$= \frac{n(n-1)R(R-1)}{N(N-1)} \sum_{y=0}^{n-2} \frac{\binom{R-2}{y} \binom{(N-2)-(R-2)}{(n-2)-y}}{\binom{N-2}{n-2}} = \frac{n(n-1)R(R-1)}{N(N-1)}$$

pmf of hypergeometric($n-2, N-2, R-2$)

$$Var(X) = E(X^2) - [E(X)]^2 = [E(X^2) - E(X)] + E(X) - [E(X)]^2$$

$$= \frac{n(n-1)R(R-1)}{N(N-1)} + \frac{nR}{N} - \left(\frac{nR}{N} \right)^2 = \frac{nR(N-R)(N-n)}{N^2(N-1)}$$

Recall.
binomial \leftrightarrow Poisson

Theorem. Let $N_i \rightarrow \infty$ and $R_i \rightarrow \infty$ in such a way that

$$p_i \equiv R_i/N_i \rightarrow p, \quad \text{cf. } n \text{ fixed}$$

where $0 < p < 1$, then

$$n \ll N_i$$

Intuition: When # of red & white balls are very large, n relatively small, without replacement \approx with replacement

pmf of hypergeometric (dep. case)

$$\frac{\binom{R_i}{k} \binom{N_i - R_i}{n-k}}{\binom{N_i}{n}} \rightarrow \binom{n}{k} p^k (1-p)^{n-k}$$

pmf of binomial (indep. case)

proof.

$$\frac{\binom{R_i}{k} \binom{N_i - R_i}{n - k}}{\binom{N_i}{n}} = \frac{R_i!}{k! (R_i - k)!} \cdot \frac{(N_i - R_i)!}{(n - k)! [(N_i - R_i) - (n - k)]!} \cdot \frac{n! (N_i - n)!}{N_i!}$$

$$= \frac{n!}{k! (n - k)!} \cdot \frac{R_i \times (R_i - 1) \times \dots \times (R_i - k + 1)}{N_i \times (N_i - 1) \times \dots \times (N_i - n + 1)} \cdot \frac{N_i \times (N_i - 1) \times \dots \times (N_i - n + 1)}{N_i \times (N_i - 1) \times \dots \times (N_i - n + 1)}$$

$$\rightarrow \binom{n}{k} p^k (1 - p)^{n - k}$$

➤ Summary for $X \sim \text{Hypergeometric}(n, N, R)$

- Range: $\mathcal{X} = \{0, 1, 2, \dots, n\} \rightarrow \max(0, n + R - N) \leq X \leq \min(R, n)$
- Pmf: $f_X(x) = \binom{R}{x} \binom{N - R}{n - x} / \binom{N}{n}$, for $x \in \mathcal{X}$
- Parameters: $n, N, R \in \{1, 2, 3, \dots\}$ and $n \leq N, R \leq N$
- Mean: $E(X) = nR/N$
- Variance: $\text{Var}(X) = nR(N - R)(N - n) / (N^2(N - 1))$

❖ Reading: textbook, Sec 4.6, 4.7, 4.8.1~4.8.3

