

Ans: 3 commonly used tools to define the p.m.'s of discrete r.v.'s:

others:
• hazard
function
• characteristic
function

also
define
for
continuous
r.v.'s

1. Probability mass function (pmf)
2. Cumulative distribution function (cdf)
3. Moment generating function (mgf, Chapter 7)

When any one of them is known, the remaining 2 can be derived from it.

- Definition: If X is a discrete r.v., then the probability mass function of X is defined by

$$f_X: \mathbb{R} \rightarrow [0, 1] \quad f_X(x) \equiv P_X(\{X = x\}) = P(\{\omega \in \Omega : X(\omega) = x\})$$

for $x \in \mathbb{R}$. (c.f., the $p: \Omega \rightarrow [0, 1]$ in LNp.3-7)

$P: 2^\Omega \rightarrow [0, 1]$ discrete sample space

Recall. For discrete sample space, it's enough to define small p .

$\because f_X(x) = 0$ for $x \notin \mathcal{X}$. It's enough to evaluate $f_X(x)$ for $x \in \mathcal{X}$

Example. For the X_1 in the Coin Tossing example, LNp.5-3

$$\mathcal{X} = \{0, 1, 2, 3\}$$

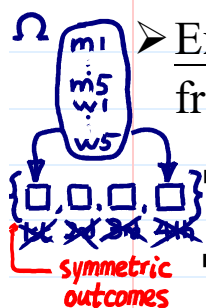
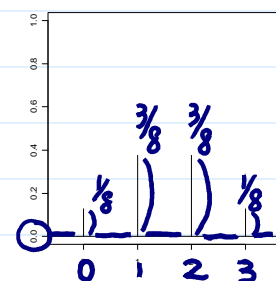
$$f_{X_1}(0) = 1/8, \quad f_{X_1}(1) = 3/8, \\ f_{X_1}(2) = 3/8, \quad f_{X_1}(3) = 1/8.$$

the small p in LNp.3-7

$$p: \Omega \rightarrow [0, 1] \\ f_{X_1}: \mathcal{X} \rightarrow [0, 1]$$

$$\text{and } f_{X_1}(x) = 0, \text{ for } x \notin \mathcal{X}.$$

- Graphical display



Example (Committees). A committee of size $n=4$ is selected from 5 men and 5 women. Then,

Ω is finite & symmetric outcomes

$$\Omega = \{\text{combination of 4}\}, \quad \#\Omega = \binom{10}{4} = 210, \quad P(A) = \#A / \#\Omega$$

Let X be the number of women on the committee, then

$$f_X(x) = P_X(X = x) = \frac{\binom{5}{x} \binom{5}{4-x}}{\binom{10}{4}}$$

$$f_X(0) = f_X(4) = \frac{5}{210}, \quad f_X(1) = f_X(3) = \frac{50}{210},$$

$$f_X(2) = \frac{100}{210}. \quad \mathcal{X} = \{0, 1, \dots, 4\}, \quad f_X(x) = 0 \text{ if } x \notin \mathcal{X}.$$

Q: What should a pmf look like?

discrete

a countably many set

Theorem. If f_X is the pmf of r.v. X with range $\mathcal{X}$, then

There exist only finite or countably infinite $x \in \mathbb{R}$ s.t. $f_X(x) > 0$

$$(i) \quad f_X(x) \geq 0, \text{ for all } x \in \mathbb{R},$$

$$(ii) \quad f_X(x) = 0, \text{ for } x \notin \mathcal{X},$$

$$(iii) \quad \sum_{x \in \mathcal{X}} f_X(x) = 1.$$

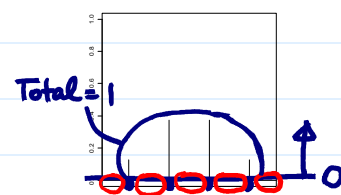
$$(iv) \quad \text{moreover, for } A \subset \mathbb{R},$$

$$P_X(X \in A) = \sum_{x \in A \cap \mathcal{X}} f_X(x).$$

can be an uncountable set

extension from small p to large P (LNp.3-7)

countably many sum



To define P_X , it's enough to define $f_X: \mathbb{R} \rightarrow \mathbb{R}$

proof (i) & (ii) holds by the definition of pmf.

Actually, $(\because f_X(x) = P(\{\omega \in \Omega \mid X(\omega) = x\}))$

$P_X(\{x \mid f_X(x) = 0\})$

= 0, for discrete r.v.'s

not true for continuous r.v.'s

check uniform spinner example in Lnp. 3-18

(iv) For $A \subset \mathbb{R}$, let $A \cap \mathcal{X} = \{x'_1, x'_2, x'_3, \dots\}$

$$P_X(A) = P(\{\omega \in \Omega \mid X(\omega) \in A \cap \mathcal{X}\}) + P(\{\omega \in \Omega \mid X(\omega) \in A \cap \mathcal{X}^c\}) = \phi$$

$$P_X(A \cap \mathcal{X}) + P_X(A \cap \mathcal{X}^c) = \sum_{k=1}^{\infty} P(\{\omega \in \Omega \mid X(\omega) = x'_k\}) = \sum_{k=1}^{\infty} f_X(x'_k) = \sum_{x \in A \cap \mathcal{X}} f_X(x)$$

(iii) follows (iv) by letting the A in (iv) be $\mathcal{X}$, then

$$\sum_{x \in \mathcal{X}} f_X(x) = P_X(\mathcal{X}) = P(\Omega) = 1.$$

Theorem. Any function f that satisfies (i), (ii), and (iii) for some finite or countably infinite set $\mathcal{X}$ is the pmf of some discrete random variable $X: \Omega \rightarrow \mathbb{R}$ $\hookrightarrow \mathbb{C} \subset \mathbb{R}$

proof. For given $\mathcal{X}$ & f , let $\Omega = \mathcal{X} \subset \mathbb{R}$.

For any $A \subset \mathcal{X}$, let $P(A) \equiv \sum_{x \in A} f(x)$. — (*)

Then, $(\Omega, \{A\} = 2^\Omega, P)$ is a probability space (exercise)

Let $X: \Omega \rightarrow \mathbb{R}$ s.t. $X(\omega) = \omega, \forall \omega \in \Omega$.

Then, X is a discrete random variable with distribution P_X , and for any $x \in \mathbb{R}$,

$$f_X(x) = P_X(\{x\}) = P(\{x\}) \stackrel{\text{by (*)}}{=} f(x).$$

"sort of" ignore the original probability space $(\Omega, \mathcal{F}, P)$

Henceforth, we can define "pmf" as any function that satisfies (i), (ii), and (iii).

We can specify a distribution by giving $\mathcal{X}$ and f , subject to the three conditions (i), (ii), (iii).

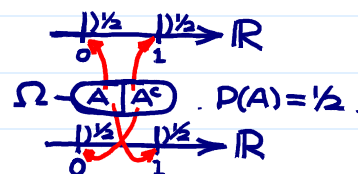
Q: Suppose that X and Y are two r.v.'s defined on Ω with the same pmf. Is it

Ans. No. always true that $X(\omega) = Y(\omega)$ for $\omega \in \Omega$?

i.e. same r.v.

Definition: A function $F_X: \mathbb{R} \rightarrow \mathbb{R}$ is called the cumulative distribution function of a random variable X if $F_X(x) = P_X(X \leq x), x \in \mathbb{R}$.

(Note. The definition of cdf can be applied to arbitrary r.v.'s) $\leftarrow$ not only applied on discrete r.v.'s. $\leftarrow$ pmf defined only for $\mathcal{X}$



(*) $\mathcal{X} = \{x, x^a, x^b\}$

① F_X define P_X on $(-\infty, x], \forall x \in \mathbb{R}$

② extend P_X to $(a, b], -\infty < a < b < \infty$
 $(\because (a, b] = (-\infty, b] - (-\infty, a])$
 $P_X(a, b] = P_X((-\infty, b]) - P_X((-\infty, a])$)

③ extend P_X to any (measurable) sets (check example in Lnp. 3-18).

Example. For the X_1 in the Coin Tossing example,

Why is it called "cumulative"?

$$F_{X_1}(x) = \begin{cases} 0, & x < 0, \\ 1/8, & 0 \leq x < 1, \\ 4/8, & 1 \leq x < 2, \\ 7/8, & 2 \leq x < 3, \\ 1, & 3 \leq x. \end{cases}$$

