

機率論 hw6

5.1

(a)

$$\begin{aligned} \int_2^4 c\left(x - \frac{3}{x^2}\right) dx &= 1 \\ \Rightarrow \left[\frac{1}{2}cx^2 + \frac{3c}{x} \right]_2^4 &= 8c + \frac{3}{4}c - 2c - \frac{3}{2}c = \frac{21}{4}c = 1 \\ \Rightarrow c &= \frac{4}{21} \end{aligned}$$

(b)

$$\begin{aligned} F_X(x) &= \begin{cases} 0, & \text{if } x < 2 \\ \int_2^x \frac{4}{21} \left(t - \frac{3}{t^2}\right) dt, & \text{if } 2 \leq x < 4 \\ 1, & \text{if } x \geq 4 \end{cases} \\ &= \int_2^x \frac{4}{21} \left(t - \frac{3}{t^2}\right) dt \\ &= \left[\frac{2}{21}t^2 + \frac{4}{7t} \right]_2^x \\ &= \frac{2}{21}x^2 + \frac{4}{7x} - \frac{8}{21} - \frac{2}{7} = \frac{2x^3 - 14x + 12}{21x} \\ F_X(x) &= \begin{cases} 0, & \text{if } x < 2 \\ \frac{2x^3 - 14x + 12}{21x}, & \text{if } 2 \leq x < 4 \\ 1, & \text{if } x \geq 4 \end{cases} \end{aligned}$$

5.4

(a)

$$P(X > 20) = \int_{20}^{\infty} \frac{10}{x^2} dx = \left[10\left(-\frac{1}{x}\right) \right]_{20}^{\infty} = \frac{1}{2}$$

(b)

$$F_X(x) = P(X \leq x) = \begin{cases} 0, & x < 10 \\ \int_{10}^x \frac{10}{t^2} dt = \left[10(-\frac{1}{t}) \right]_{10}^x = 1 - \frac{10}{x}, & x \geq 10 \end{cases}$$

(c)

Assumption: 每個裝置壽命 > 15 的事件彼此獨立

$$P(X > 15) = 1 - F_X(15) = 1 - (1 - \frac{10}{15}) = \frac{2}{3}$$

令 Y 為這 6 個裝置中壽命 > 15 的裝置個數，則 $Y \sim \text{binomial}(6, \frac{2}{3})$

$$P(Y \geq 3) = \sum_{i=3}^6 \binom{6}{i} P(X > 15)^i P(X \leq 15)^{6-i} = \sum_{i=3}^6 \binom{6}{i} \left(\frac{2}{3}\right)^i \left(\frac{1}{3}\right)^{6-i}$$

5.6

(a)

$$E(X) = \int_0^\infty \frac{1}{4} x^2 e^{-\frac{x}{2}} dx = \int_0^\infty 2y^2 e^{-y} dy = 2\Gamma(3) = 2E(Y^2), \quad Y \sim \text{exp}(1)$$

so

$$E(X) = 2E(Y^2) = 2\left(Var(Y) + (E(Y))^2\right) = 2 \times 2 = 4$$

另一種算法是直接使用 Gamma function: $2 \times \Gamma(3) = 2 \times 2! = 4$

(b)

$$\int_{-1}^1 c(1-x^2)dx = \left[cx - \frac{1}{3}cx^3 \right]_{-1}^1 = c - \frac{1}{3}c + c - \frac{1}{3}c = \frac{4}{3}c = 1$$

$$\Rightarrow c = \frac{3}{4}$$

$$E(X) = \int_{-1}^1 \frac{3}{4}x(1-x^2)dx = \int_{-1}^1 \frac{3}{4}x - \frac{3}{4}x^3 dx = \left[\frac{3}{8}x^2 - \frac{3}{16}x^4 \right]_{-1}^1 = 0$$

也可直接看出 pdf 是偶函數 i.e., $f(x)=f(-x)$, 所以:

$$\int_{-1}^1 xf(x)dx$$

令 $y = -x$

$$= \int_1^{-1} yf(-y)dy = -\int_{-1}^1 yf(-y)dy \Rightarrow 2 \int_{-1}^1 xf(x)dx = 0 \Rightarrow \int_{-1}^1 xf(x)dx = 0$$

(c)

$$E(X) = \int_5^\infty x \frac{5}{x^2} dx = \int_5^\infty \frac{5}{x} dx = \left[5 \ln x \right]_5^\infty = \infty \Leftrightarrow E(X) \text{ 不存在}$$

5.11

Def:

X : 在長度為 L 的線段上選取一點的位置, $0 \leq X \leq L$, 因為隨機選取, 故

$$X \sim \text{Unif}(0, L)$$



$$\begin{aligned} P\left(\frac{\text{the shorter segment}}{\text{the longer segment}} < \frac{1}{4}\right) &= P\left(\min\left(\frac{X}{L-X}, \frac{L-X}{X}\right) < \frac{1}{4}\right) = P\left(\frac{X}{L-X} < \frac{1}{4} \text{ or } \frac{L-X}{X} < \frac{1}{4}\right) \\ &= P\left(X < \frac{L}{5} \text{ or } X > \frac{4L}{5}\right) = \int_0^{\frac{L}{5}} \frac{1}{L} dx + \int_{\frac{4L}{5}}^L \frac{1}{L} dx \\ &= \frac{1}{5} + \frac{1}{5} = \frac{2}{5} \end{aligned}$$

5.13

(a)

$X \equiv$ passenger waiting time, 則 $0 \leq X \leq 15$ 又因 X 服從 uniform, 故 $X \sim U(0, 15)$

$$P(X > 6) = \int_6^{15} \frac{1}{15} dx = \frac{9}{15} = \frac{3}{5}$$

(b)

$$P(X > 10 | X > 8) = \frac{P(X > 10, X > 8)}{P(X > 8)} = \frac{P(X > 10)}{P(X > 8)} = \frac{\int_{10}^{15} \frac{1}{15} dx}{\int_8^{15} \frac{1}{15} dx} = \frac{5}{7}$$

5.38

(a)

$$f_X(x) = \frac{1}{4}, \quad 1 < x < 5$$

因 $y = \ln(x)$ 是 strictly monotone function on $1 < x < 5$, 故

$$y = \ln x \Rightarrow x = e^y \Rightarrow \frac{dx}{dy} = e^y$$

只是提醒公式裡要加絕對值,

雖然這題不會有影響

$$f_Y(y) = f_X(e^y) \cdot |\widehat{e^y}| = \frac{1}{4} e^y, \quad 0 < y < \ln 5, \text{ and } f_Y(y) = 0, \text{ otherwise}$$

(b)

$$P\left(\frac{1}{3} < Y < \frac{2}{3}\right) = \int_{\frac{1}{3}}^{\frac{2}{3}} \frac{1}{4} e^y dy = \left[\frac{1}{4} e^y\right]_{\frac{1}{3}}^{\frac{2}{3}} = \frac{1}{4} e^{\frac{2}{3}} - \frac{1}{4} e^{\frac{1}{3}}$$

5.42

For a set $A \subset \mathbb{R}$, define the indicator function of A as $I_A(x) = \begin{cases} 0, & \text{if } x \notin A \\ 1, & \text{if } x \in A \end{cases}$.

The pdf of X is $f_X(x) = \frac{1}{b-a} I_{(a,b)}(x)$.

First note that $Y = d$, if $c = 0$ (i.e., $\Pr(Y = d) = 1$) $\forall d \in \mathbb{R}$.

Now suppose $c \neq 0 \implies Y = cX + d$ is a strictly monotone function of x and $x = \frac{y-d}{c} \equiv g^{-1}(y)$

We have $f_Y(y) = f_X\left(\frac{y-d}{c}\right) \times |J|$, where

$$|J| = \left| \frac{dg^{-1}(y)}{dy} \right| = \left| \frac{dx}{dy} \right| = \begin{cases} \frac{1}{c}, & \text{if } c > 0 \\ -\frac{1}{c}, & \text{if } c < 0 \end{cases}$$

$$So f_Y(y) = \begin{cases} \frac{1}{c(b-a)} I_{(ac+d, bc+d)}(y), & if c > 0. \\ \frac{1}{-c(b-a)} I_{(bc+d, ac+d)}(y), & if c < 0. \end{cases}$$

$$That is, Y \begin{cases} \sim U(ac + d, bc + d), & if c > 0 \\ = d, & if c = 0 \\ \sim U(bc + d, ac + d), & if c < 0 \end{cases}$$

TE2

$$E[Y] = \int_0^\infty P(Y > y) dy - \int_0^\infty P(Y < -y) dy$$

$$\int_0^\infty P(Y < -y) dy = \int_0^\infty \int_{-\infty}^{-y} f_Y(x) dx dy = \int_{-\infty}^0 \int_0^{-x} f_Y(x) dy dx$$

$$= \int_{-\infty}^0 (-x) f_Y(x) dx = - \int_{-\infty}^0 x f_Y(x) dx = - \int_{-\infty}^0 y f_Y(y) dy$$

$$\int_0^\infty P(Y > y) dy = \int_0^\infty \int_y^\infty f_Y(x) dx dy$$

$$= \int_0^\infty \int_0^x f_Y(x) dy dx$$

$$= \int_0^\infty x f_Y(x) dx$$

$$= \int_0^\infty y f_Y(y) dy$$

$$\therefore E[Y] = \int_0^\infty y f_Y(y) dy + \int_{-\infty}^0 y f_Y(y) dy$$

$$= \int_0^\infty y f_Y(y) dy - \left[- \int_{-\infty}^0 y f_Y(y) dy \right]$$

$$= \int_0^\infty P(Y > y) dy - \int_0^\infty P(Y < -y) dy$$

TE5

$$\begin{aligned}
 E[X^n] &= \int_0^\infty P(X^n > t) dt \\
 &= \int_0^\infty nx^{n-1} P(X^n > x^n) dx \quad (*) \\
 &= \int_0^\infty nx^{n-1} P(X > x) dx \\
 (*) : \text{let } t &= x^n, dt = nx^{n-1}dx
 \end{aligned}$$

TE8

$$0 \leq X \leq c \Rightarrow 0 \leq X^2 \leq cX$$

$$\Rightarrow 0 \leq E(X^2) \leq cE(X)$$

$$\text{Var}(X) = E(X^2) - [E(X)]^2$$

$$\leq cE(X) - [E(X)]^2$$

$$= c \cdot (c\alpha) - (c\alpha)^2, \quad \text{where } \alpha = \frac{E(X)}{c}$$

$$= c^2\alpha(1 - \alpha)$$

$$= -c^2 \left(\alpha - \frac{1}{2}\right)^2 + \frac{1}{4}c^2$$

$$\begin{aligned}
 &\leq \underbrace{\frac{1}{4}c^2}_{\text{提醒若 } X = \begin{cases} c, & \text{with prob. } \frac{1}{2} \\ 0, & \text{with prob. } \frac{1}{2} \end{cases}, \text{ 則等號成立}}
 \end{aligned}$$